Final Exam Probability and Statistics EE1M31 18 April 2018, 9.00 – 11.00

Last name, initials:	Grade:
Student number:	

Open questions

ALWAYS MOTIVATE YOUR ANSWERS AND WRITE NEATLY.

YOU MAY ANSWER IN EITHER DUTCH OR ENGLISH.

1. The lifetime X of an electrical component is a continuous random variable having the probability density function $f_{\theta}(x) = \frac{2}{\theta} - \frac{2}{\theta^2}x$ for $0 \le x \le \theta$ and $f_{\theta}(x) = 0$ otherwise.

Let X_1, \ldots, X_n be a random sample from f_{θ} .

(a) (3 points) Investigate whether $S = \sqrt{\frac{6}{n}(X_1^2 + \ldots + X_n^2)}$ is an unbiased estimator for θ . Hint: first investigate whether S^2 is an unbiased estimator for θ^2 .

$$E[X_{1}^{2}] = \int_{0}^{1} x^{2} \left(\frac{2}{\theta} - \frac{2}{\theta^{2}}x\right) dx = \frac{2}{\theta} \int_{0}^{1} x^{2} dx - \frac{2}{\theta^{2}} \int_{0}^{1} x^{3} dx$$

$$= \frac{2}{\theta} \left[\frac{1}{3}x^{3}\right]_{0}^{0} - \frac{2}{\theta^{2}} \left[\frac{1}{4}x^{4}\right]_{0}^{0} = \frac{2}{3}\theta^{2} - \frac{2}{\theta^{2}} \int_{0}^{1} x^{3} dx$$

$$= \left(\frac{2}{3} - \frac{1}{2}\right)\theta^{2} = \left(\frac{4}{6} - \frac{3}{6}\right)\theta^{2} = \frac{1}{6}\theta^{2}$$

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$$= \left(\frac{1}{3} + \frac{3}{6}\right)\theta^{2} = \frac{1}{6}\theta^{2}$$

$$= \frac{1}{6}\theta^{2} + \frac{3}{6}\theta^{2} + \frac{3}{6}\theta^{2} = \frac{1}{6}\theta^{2}$$

$$= \frac{1$$

(b) (3 points) Let $T = \frac{3}{n}(X_1 + \ldots + X_n)$ be another estimator for θ : determine its MSE.

$$E[X_{n}] = \int_{0}^{\infty} x \left(\frac{2}{6} - \frac{2}{6^{2}}x\right) dx = \frac{2}{6} \int_{0}^{\infty} x dx - \frac{2}{6^{2}} \int_{0}^{\infty} x^{2} dx$$

$$= \frac{2}{6} \frac{1}{2} 0^{2} - \frac{2}{6^{2}} \frac{1}{3} 0^{3} = 0 - \frac{2}{3} 0 = \frac{1}{3} 0$$

$$E[T] = \frac{3}{5} \cdot n \cdot \frac{1}{3} 0 = 0 = 0 \quad \text{on sinsed}$$

$$MSE[T] = Var(T) = \frac{9}{5} \cdot n \cdot Var(X_{n}) = \frac{9}{5} \cdot n \cdot E[X_{n}] - E[X_{n}]^{2}$$

$$= \frac{9}{5} \cdot \left(\frac{1}{6} 0^{2} - \frac{1}{3} 0^{2}\right) = \frac{1}{5} \cdot \left(15 0^{2} - 0^{2}\right) = \frac{0^{2}}{2} \cdot n$$

2. Consider the following ordered dataset:

$$3.14 \quad 3.22 \quad 3.43 \quad 3.73 \quad 3.94 \quad 3.96 \quad 4.06 \quad 4.22 \quad 4.22 \quad 5.60 \quad 7.49$$

(a) (2 points) Calculate the 20th empirical percentile of this dataset.

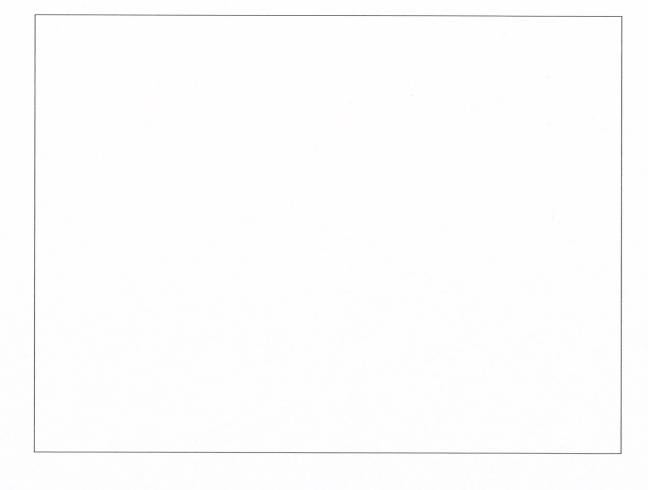
$$p=0.2$$
 $n=11$ $p(n+1)=0.2$. $12=2.4$ $k=[p(n+1)]=2$ $x=p(n+1)-K=0.4$ (see p. 235 in the stree book)

 $q_{11}(0.2)=3.22+0.4$ (3.43-3.22) = 3.22+0.4 .0.21 = 3.304

Assume that the dataset above has been sampled from the following density:

$$f_{\theta}(x) = \begin{cases} 0 & x < 3, \\ \frac{1}{\theta} e^{-\frac{(x-3)}{\theta}} & x \ge 3. \end{cases}$$

(b) (4 points) Calculate the maximum likelihood estimate for θ . You may use that $\sum x_i = 47.01$, $\sum \ln x_i = 15.63$, $\sum \frac{1}{x_i} = 2.73$ or $\sum e^{-x_i} = 0.23$.



- 3. A manufacturer produces two new types of solar panels: type A and type B. The peak power of type A has a normal distribution with unknown expectation μ_A , but the variance is known: $\sigma_A^2 = 100$. The manufacturer wishes to estimate μ_A and therefore measures the peak power of 20 solar panels of type A.
 - (a) (3 points) Explain with a computation why $\bar{x}_{20} \pm 1.96 \frac{10}{\sqrt{20}}$ is the two-sided 95% confidence interval for μ_A .

$$\overline{X}_{n}-\mu$$

is distributed as $N(0,1)$ since \overline{X}_{n} is

 $\overline{O}_{A}/\sqrt{n}$ an average of normal random \sqrt{no} ; ables

 $0.95 = P(-\overline{Z}_{0.025} \subseteq \overline{X}_{A} - \mu) \subseteq \overline{Z}_{0.025})$
 $= P(-\overline{Z}_{0.025} \subseteq \overline{X}_{A} \subseteq \overline{X}_{n} - \mu) \subseteq \overline{Z}_{0.025} \subseteq \overline{X}_{n})$
 $= P(-\overline{Z}_{0.025} \subseteq \overline{X}_{A} \subseteq \mu - \overline{X}_{n} \subseteq \overline{Z}_{0.025} \subseteq \overline{X}_{n})$
 $= P(\overline{X}_{n} - \overline{Z}_{0.025} \subseteq \overline{X}_{n} \subseteq \mu \subseteq \overline{X}_{n} + \overline{Z}_{0.025} \subseteq \overline{X}_{n})$

We see that $(\overline{X}_{20} - 1.95 \stackrel{10}{\sqrt{10}} + \overline{X}_{20} + 1.95 \stackrel{10}{\sqrt{20}})$

is a 95% confidence interval.

For type B, the manufacturer claims that $\mu_B = 300\,W$. However, the Consumers Association received some negative feedback that the actual peak power is lower so they decided to investigate this claim. They tested 15 panels of type B and found a sample mean of 290.3 W and a sample variance of 353.8 W^2 . You may assume a normal distribution as model distribution.

(b) (3 points) Investigate the claim of the manufacturer at significance level $\alpha = 0.10$.

We test to: MB = 300 against this Mg < 300, so me have a one-sided test.

$$T = \frac{x_{15} - 300}{\sqrt{353.8} / \sqrt{45}} = \frac{290.3 - 300}{\sqrt{353.8} / \sqrt{45}} = -1.997$$
The critical value is $-t_{14}$, on $=-1.345$.

$$-1.997c - 1.345$$
, so we reject to.

The test yields that the peaf power is lower than the claimed. $300W$.

Answers multiple choice:

- e.
- c.
- f.
- a.
- d.
- a.
- d.
- 8 a.
- 9 c.